

Filtered Markovian Projection: Dimensionality Reduction in Filtering for Stochastic Reaction Networks

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NDNS+ WORKSHOP 2025, University of Twente, The Netherlands

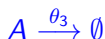
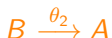
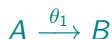
June 24, 2025

Stochastic Reaction Networks (SRNs): Motivation

Aim: Model the time evolution of populations (molecules, species, agents, cells, ...) connected through reaction networks.

Illustration:

Reaction Network involving
molecules/Species/Agents:
A and B



$$\begin{pmatrix} 0 \\ 0 \end{pmatrix} \quad \begin{pmatrix} 1 \\ 0 \end{pmatrix} \quad \begin{pmatrix} 2 \\ 0 \end{pmatrix} \quad \dots$$

$$\begin{pmatrix} 0 \\ 1 \end{pmatrix} \quad \begin{pmatrix} 1 \\ 1 \end{pmatrix} \quad \begin{pmatrix} 2 \\ 1 \end{pmatrix} \quad \dots$$

$$\begin{pmatrix} 0 \\ 2 \end{pmatrix} \quad \begin{pmatrix} 1 \\ 2 \end{pmatrix} \quad \begin{pmatrix} 2 \\ 2 \end{pmatrix} \quad \dots$$

$$\begin{pmatrix} x_1 \\ x_2 \end{pmatrix} \begin{matrix} \# \text{ type } A \\ \# \text{ type } B \end{matrix}$$

$$\vdots$$

$$\vdots$$

$$\vdots$$

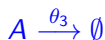
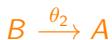
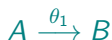
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Stochastic Reaction Networks (SRNs): Motivation

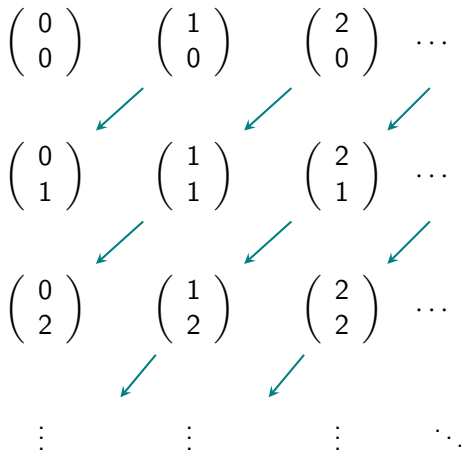
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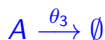
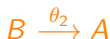
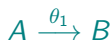


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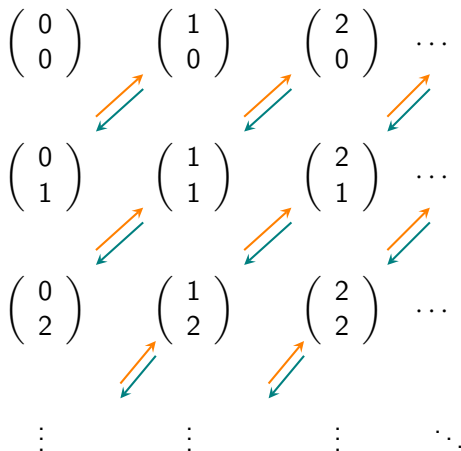
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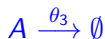
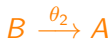
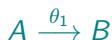


Stochastic Reaction Networks (SRNs): Motivation

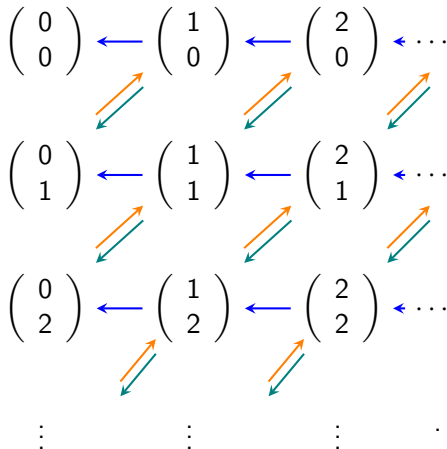
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Stochastic Reaction Networks (SRNs): Motivation

Aim: Model the time evolution of populations (molecules, species, agents, cells, ...) connected through reaction networks.

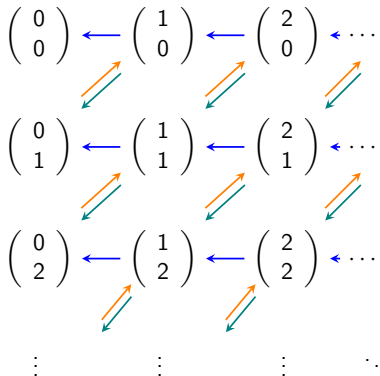
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$$\begin{pmatrix} x_1 \\ x_2 \end{pmatrix} \# \text{ type } A$$

$$\# \text{ type } B$$



⚠ For d number of molecule types/species/agents $\Rightarrow$ state space: $\mathbb{N}^d$

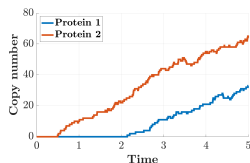
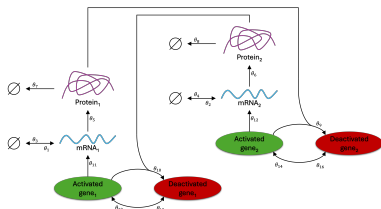
Motivation: Dimensionality Reduction

- **Challenge:** Numerical solutions for problems modeled by Stochastic Reaction Networks (SRNs) (particular type of Markov chains) often suffer from the **curse of dimensionality**
 - **Forward Kolmogorov / Chemical Master Equations** (time evolution of the joint (full) probability distribution over the possible states of the system).
 - **Stochastic Filtering** (i.e., estimate the conditional probability distribution of unobserved states for partially observed systems).
 - **Stochastic Optimal Control**.

Motivation: Dimensionality Reduction

- **Challenge:** Numerical solutions for problems (e.g., Chemical Master Equations, Stochastic Filtering, Stochastic Optimal Control) modeled by SRNs often suffer from the **curse of dimensionality**.
- **Observation:** Quantities of interest often do not depend on the full-dimensional state (e.g., $\mathbb{E}[g(\mathbf{X}(\mathbf{t}))]$ with $g(\mathbf{x}) = \mathbb{1}_{\{x_i > \gamma\}}$).

Example: Bistable Gene Switch Network



Observed trajectory of
Protein₁ and **Protein₂**.

We are interested in the distribution of the (hidden) mRNA₂ copy number.

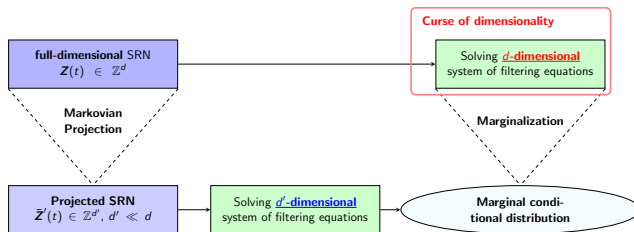
Motivation: Dimensionality Reduction

- **Challenge:** Numerical solutions for problems (e.g., Chemical Master Equations, Stochastic Filtering, Stochastic Optimal Control) modeled by SRNs often suffer from the **curse of dimensionality**.
-  Quantities of interest often do not depend on the full-dimensional state
⇒ **Solution:**
 - ① construct a low dimensional Markovian process that mimics the evolution of the projected process (preserve the marginal distribution of the original process)
 - ② ⇒ Solve the resulting reduced-dimensional problem.

Related Work

C. Ben Hammouda, M. Chupin, S Munker, and R Tempone. “Filtered Markovian Projection: Dimensionality Reduction in Filtering for Stochastic Reaction Networks”. In: *arXiv preprint arXiv:2502.07918* (2025)

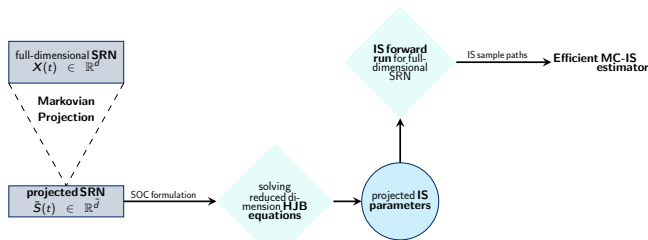
- 1 **Inverse Problem:** Design efficient **finite-dimensional stochastic filters** to estimate the **conditional probability distribution of unobserved (hidden) state variables** given partial observation trajectories.
 - Approach based on solving the system of filtering equations (an infinite-dimensional system of coupled ODEs with jumps) in a projected low-dimensional space via Markovian projection.



Related Work

C. Ben Hammouda, N. Ben Rached, R. Tempone, and S. Wiechert. “Automated importance sampling via optimal control for stochastic reaction networks: A Markovian projection–based approach”. In: *Journal of Computational and Applied Mathematics* 446 (2024), p. 115853

- ② **Forward Problem:** Design efficient Monte Carlo (MC) estimators for rare event probabilities via a generic path dependent measure change (importance sampling).

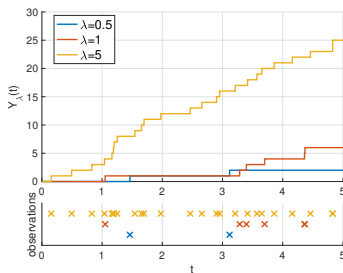


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Stochastic Reaction Networks (SRNs): Motivation

Aim: Model the evolution of populations connected through networks.

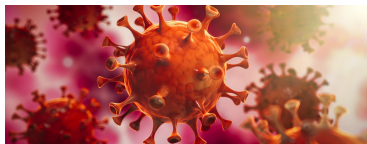
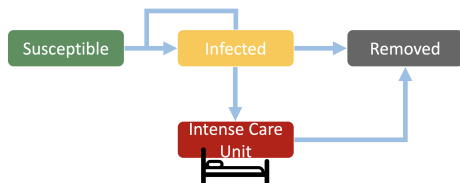
- Deterministic models (based on ODEs) capture the **macroscopic** behavior and are **only valid for large populations**.
- Presence of species/agents with **small population sizes** introduces **stochastic effects** (Elowitz et al. 2002; Raj et al. 2006; Erban et al. 2020)
⇒ **SRNs**: continuous-time, discrete-space Markov chains driven by **Poisson processes**.



- Realizations of $Y_\lambda(t) \sim \text{Poisson}(\lambda t)$.
- For $0 < s < t$: $\mathbb{P}(Y_\lambda(t) - Y_\lambda(s) = k) = \frac{(\lambda(t-s))^k}{k!} e^{-\lambda(t-s)}$
- Observations given as crosses (correspond to a jump).
- Inter arrival time of Y_λ is exponentially distributed with parameter $\lambda > 0$.

Stochastic Reaction Networks (SRNs): Applications

① Epidemiological networks (Brauer et al. 2012; Anderson et al. 2015).

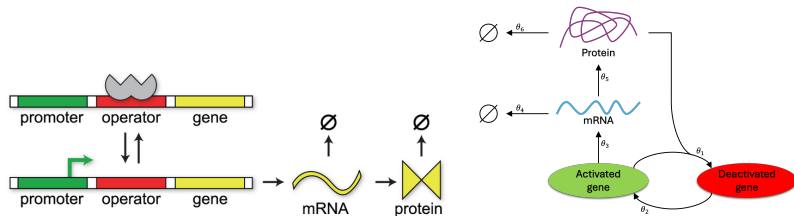


Example: SIR Model for epidemics. Removed means either recovered or dead



Stochastic Reaction Networks (SRNs): Applications

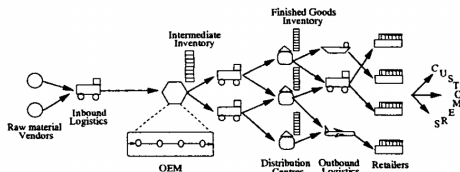
- 1 Epidemiological networks (Brauer et al. 2012; Anderson et al. 2015).
- 2 Transcription and translation in genomics and virus Kinetics (Peccoud et al. 1995; Hensel et al. 2009; Roberts et al. 2011)



Example: Simple genetic switch (regulation)

Stochastic Reaction Networks (SRNs): Applications

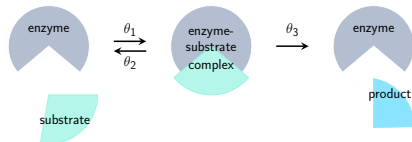
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- 3 Manufacturing supply chain, logistic/social/ecological networks, (Raghavan et al. 2002; Goutsias et al. 2013; Chiang et al. 2020)



Example:
Supply chain network

Stochastic Reaction Networks (SRNs): Applications

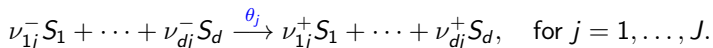
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- 4 (Bio)chemical reactions (Rao et al. 2003; Briat et al. 2015),...



Example: Michaelis-Menten enzym kinetics

Stochastic Reaction Networks: Mathematical Formulation

Consider a reaction network consisting of d species/agents $S_1, \dots, S_d$ and J reaction channels:



- A **stochastic reaction network** is a **continuous-time discrete-space Markov chain**, $\mathbf{X}(t)$, defined on a probability space $(\Omega, \mathcal{F}, \mathbb{P})$

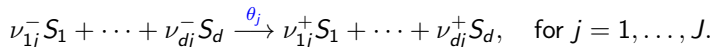
$$\mathbf{X}(t) = (X^{(1)}(t), \dots, X^{(d)}(t)) : [0, T] \times \Omega \rightarrow \mathbb{N}^d$$

- $X^{(i)}(t)$ models the counting number of the i -th agent/species at time t .
- The dynamics of $\mathbf{X}$ is characterized by $(\boldsymbol{\nu}_j, \mathbf{a}_j)_{j=1}^J$, where
 - $\boldsymbol{\nu}_j := (\nu_{1j}^+ - \nu_{1j}^-, \dots, \nu_{dj}^+ - \nu_{dj}^-)^\top \in \mathbb{Z}^d$: **stoichiometric (state change) vector**.
 - $\mathbf{a}_j : \mathbb{N}^d \rightarrow \mathbb{R}_+$: **propensity (jump intensity) function**

$$\mathbb{P}(\mathbf{X}(t + \Delta t) = \mathbf{x} + \boldsymbol{\nu}_j \mid \mathbf{X}(t) = \mathbf{x}) = \mathbf{a}_j(\mathbf{x})\Delta t + o(\Delta t)$$

Stochastic Reaction Networks (SRNs): Mathematical Formulation

A reaction network consisting of d species/agents $S_1, \dots, S_d$ and J reaction channels:



- From the *mass-action kinetic principle*:

$$a_j(\mathbf{x}) := \underbrace{\theta_j \prod_{i=1}^d \frac{x_i!}{(x_i - \nu_{ij}^-)!}}_{\substack{\text{\#combinations of molecules} \\ \text{required for reaction } j}} 1_{\{x_i \geq \nu_{ij}^-\}},$$

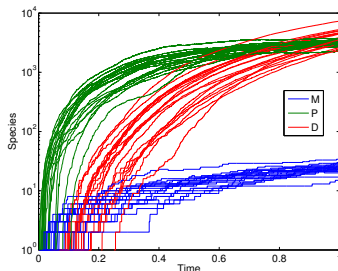
with θ_j : reaction rate of the j th reaction.

Dynamics of SRNs

- Kurtz's random time-change representation (Ethier et al. 2009)

$$\mathbf{X}(t) = \mathbf{X}(0) + \sum_{j=1}^J \underbrace{Y_j \left(\int_0^t a_j(\mathbf{X}(s)) ds \right)}_{:= R_j(t)} \cdot \boldsymbol{\nu}_j, \quad (1)$$

- $\mathbf{X}(0) \sim \mu$ is the initial state
- $\{Y_j\}_{1 \leq j \leq J}$: independent **unit-rate Poisson processes**.
- $R_j(t)$: number of occurrences of the j th reaction up to time t .



20 exact i.i.d. paths
of $X = (M, P, D)$:
counts the number of
particles of each
species in a problem
from genomics.

Dynamics of SRNs

- Kurtz's random time-change representation (Ethier et al. 2009)

$$\mathbf{X}(t) = \mathbf{X}(0) + \sum_{j=1}^J \underbrace{\gamma_j \left(\int_0^t a_j(\mathbf{X}(s)) ds \right)}_{:= R_j(t)} \cdot \nu_j, \quad (2)$$

- $\mathbf{X}(0) \sim \mu$ is the initial state.
- $\{\gamma_j\}_{1 \leq j \leq J}$: independent unit-rate Poisson processes.
- $R_j(t)$: number of occurrences of the j th reaction up to time t .
- **Non-explosivity Assumption:** Condition (A.1) ensures that the system state will almost surely not grow to infinity in a finite time interval $[0, T]$.

$$\sup_{s \in [0, T]} \sum_{j=1}^J \mathbb{E} [a_j^2(\mathbf{X}(s))] < \infty, \quad \forall T > 0 \quad (\text{A.1})$$

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Motivation

A typical challenge arising in practical problems modeled by SRNs: a few state variables can be dynamically observed $\Rightarrow$ Partially observed SRNs.

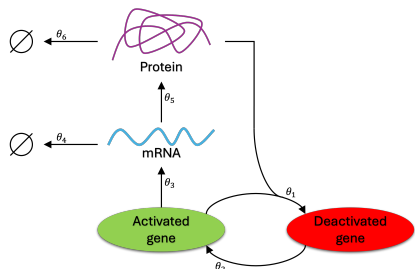


Figure: Gene Switch Network Example (Duso et al. 2018)

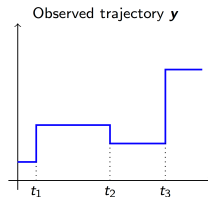
- In a real experiment, one can only measure the amount of **fluorescent protein**.
- All other hidden dynamic states of pivotal interest remain unknown, e.g.,
 - the number of **mRNA** molecules
 - the state of the gene (**on/off**)

Problem of Interest: Given partial measurement trajectories, how to efficiently estimate the conditional probability distribution of unobserved (hidden) state-variables, i.e. efficiently solve the related stochastic filtering problem?

Filtering Problem

For $t \in [0, T]$ and initial distribution $\mathbf{Z}(0) \sim \mu$

$$\mathbf{Z}(t) = \begin{pmatrix} \mathbf{X}(t) \\ \mathbf{Y}(t) \end{pmatrix} \quad \begin{array}{l} \text{hidden part} \\ \text{observed part} \end{array}$$



- **Hidden process:**

$$\mathbf{X}(t) = \mathbf{X}(0) + \sum_{j=1}^J R_j \left(\int_0^t a_j(\mathbf{X}(s), \mathbf{Y}(s)) ds \right) \boldsymbol{\nu}_{\mathbf{x},j}, \text{ where}$$

- $\boldsymbol{\nu}_{\mathbf{x},j}$ is the stoichiometric vector corresponding to $\mathbf{X}(t)$
- R_j are independent unit-rate Poisson processes

- **Observed process:**

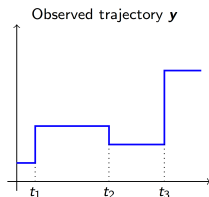
$$\mathbf{Y}(t) = \mathbf{Y}(0) + \sum_{j=1}^J R_j \left(\int_0^t a_j(\mathbf{X}(s), \mathbf{Y}(s)) ds \right) \boldsymbol{\nu}_{\mathbf{y},j}, \text{ where}$$

- $\boldsymbol{\nu}_{\mathbf{y},j}$ is the stoichiometric vector corresponding to $\mathbf{Y}(t)$

Filtering Problem

For $t \in [0, T]$ and initial distribution $\mathbf{Z}(0) \sim \mu$

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- **Hidden process:**

$$\mathbf{X}(t) = \mathbf{X}(0) + \sum_{j=1}^J R_j \left(\int_0^t a_j(\mathbf{X}(s), \mathbf{Y}(s)) ds \right) \nu_{\mathbf{x},j}$$

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Goal: Estimate conditional distribution of the hidden species/agents

$$\pi^{\mathbf{y}}(\mathbf{x}, t) := \mathbb{P}_{\mu}(\mathbf{X}(t) = \mathbf{x} \mid \mathbf{Y}(s) = \mathbf{y}(s), 0 \leq s \leq t).^1$$

¹The well-posedness of $\pi^{\mathbf{y}}(\mathbf{x}, t)$ is guaranteed by Theorem 2.24 in (Bain et al. 2009)

Filtering Problem: Setting

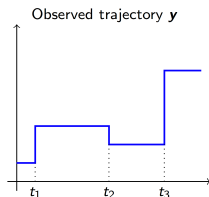
- We consider the case of noise-free and continuous in time observations.
- We assume that the trajectory $\mathbf{y}([0, T])$ can be observed with exact jump times $t_1, t_2, \dots, t_n$.
- The **observed reactions**

$$\mathcal{O}_k := \{j \in \{1, \dots, J\} : \nu_{\mathbf{y},j} = \mathbf{y}(t_k) - \mathbf{y}(t_k^-)\}$$

can only occur at $t = t_k$.

- Between jump times there could only occur **unobserved reactions**

$$\mathcal{U} := \{j \in \{1, \dots, J\} : \nu_{\mathbf{y},j} = \mathbf{0}\}$$



Filtering Equations (Zechner et al. 2014)²

$$\pi^{\mathbf{y}}(\mathbf{x}, t) := \mathbb{P}_{\mu}(\mathbf{X}(t) = \mathbf{x} \mid \mathbf{Y}(s) = \mathbf{y}(s), 0 \leq s \leq t).$$

The unnormalized conditional pmf $\rho^{\mathbf{y}}(\mathbf{x}, t) \propto \pi^{\mathbf{y}}(\mathbf{x}, t)$ satisfies

- **between jump times** for $t \in (t_k, t_{k+1})$,

$$\frac{d}{dt} \rho^{\mathbf{y}}(\mathbf{x}, t) = \sum_{j \in \mathcal{U}} a_j(\mathbf{x} - \boldsymbol{\nu}_{\mathbf{x}, j}, \mathbf{y}(t_k)) \rho^{\mathbf{y}}(\mathbf{x} - \boldsymbol{\nu}_{\mathbf{x}, j}, t) - \sum_{j=1}^J a_j(\mathbf{x}, \mathbf{y}(t_k)) \rho^{\mathbf{y}}(\mathbf{x}, t)$$

- and **at jump times** t_k

$$\rho^{\mathbf{y}}(\mathbf{x}, t_k) = \frac{1}{|\mathcal{O}_k|} \sum_{j \in \mathcal{O}_k} a_j(\mathbf{x} - \boldsymbol{\nu}_{\mathbf{x}, j}, \mathbf{y}(t_{k-1})) \rho^{\mathbf{y}}(\mathbf{x} - \boldsymbol{\nu}_{\mathbf{x}, j}, t_k^-).$$

²Existence and uniqueness of the filtering equation are shown in (D'Ambrosio et al. 2022).

Filtering Equations (Zechner et al. 2014)²

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The normalized pmf can be attained by

$$\pi^{\mathbf{y}}(\mathbf{x}, t) = \frac{\rho^{\mathbf{y}}(\mathbf{x}, t)}{\sum_{\hat{\mathbf{x}} \in X} \rho^{\mathbf{y}}(\hat{\mathbf{x}}, t)}.$$

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$$\frac{d}{dt} \rho^{\mathbf{y}}(\mathbf{x}, t) = \sum_{j \in \mathcal{U}} a_j(\mathbf{x} - \boldsymbol{\nu}_{\mathbf{x}, j}, \mathbf{y}(t_k)) \rho^{\mathbf{y}}(\mathbf{x} - \boldsymbol{\nu}_{\mathbf{x}, j}, t) - \sum_{j=1}^J a_j(\mathbf{x}, \mathbf{y}(t_k)) \rho^{\mathbf{y}}(\mathbf{x}, t)$$

- and **at jump times** t_k

$$\rho^{\mathbf{y}}(\mathbf{x}, t_k) = \frac{1}{|\mathcal{O}_k|} \sum_{j \in \mathcal{O}_k} a_j(\mathbf{x} - \boldsymbol{\nu}_{\mathbf{x}, j}, \mathbf{y}(t_{k-1})) \rho^{\mathbf{y}}(\mathbf{x} - \boldsymbol{\nu}_{\mathbf{x}, j}, t_k^-).$$

Let $\rho^{\mathbf{y}}(t) := (\rho^{\mathbf{y}}(\mathbf{x}_1, t), \rho^{\mathbf{y}}(\mathbf{x}_2, t), \dots)^{\top}$, then the filtering equations can be written as (**A** and **B** are infinite-dimensional matrices)

$$\begin{aligned} \frac{d}{dt} \rho^{\mathbf{y}}(t) &= \mathbf{A} \rho^{\mathbf{y}}(t) \quad t \in (t_k, t_{k+1}) \\ \rho^{\mathbf{y}}(t_k) &= \mathbf{B} \rho^{\mathbf{y}}(t_k^-), \end{aligned}$$

Filtered Finite State Projection (FFSP)

- Extension of **Finite State Projection** (Munsky et al. 2006) originally introduced for efficiently solving the chemical master equation.
- **Idea** (D'Ambrosio et al. 2022):
 - Consider only for a finite subset of states $\{\mathbf{x}_1, \dots, \mathbf{x}_N\} \subseteq \mathbb{N}^{\dim(\mathbf{X})}$.
 - Truncate the filtering equations $\Rightarrow$ **system of (finite-dimensional) N linear ODEs**:

$$\begin{aligned}\frac{d}{dt} \rho_N^y(t) &= \mathbf{A}_N \rho_N^y(t) \quad t \in (t_k, t_{k+1}) \\ \rho_N^y(t_k) &= \mathbf{B}_N \rho_N^y(t_k^-),\end{aligned}$$

- Control the approximation error $\mathcal{E}_{FFSP}$ by extending the truncated state space based on a posteriori estimates.

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- Control the approximation error $\mathcal{E}_{FFSP}$ by extending the truncated state space based on a posteriori estimates.



Curse of Dimensionality

If each hidden species takes n possible states
 $\rightarrow$ the truncated system contains $N = n^{\dim(\mathbf{X})}$ equations

Particle Filter (Rathinam et al. 2021)

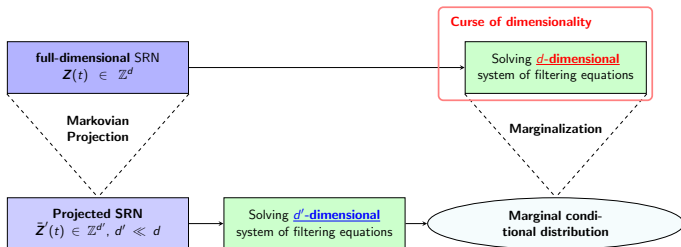
- It can be shown that: $\pi^{\mathbf{y}}(\mathbf{x}, t) \propto \mathbb{E}_{\mu} [1_{\{\mathbf{v}(t)=\mathbf{x}\}} w(t)]$, where
 - $\mathbf{V}$ is a modification of $\mathbf{X}$ in which reactions are consistent with the observed trajectory $\mathbf{y}$,
 - w is a weight that represents the importance of trajectory $\mathbf{V}(t)$.
- Based on M particles $\left\{ \hat{\mathbf{V}}_k, \hat{w}_k \right\}_{k=1, \dots, M}$, one can use the estimator

$$\pi^{\mathbf{y}}(\mathbf{x}, t) \approx \frac{1}{\sum_{k=1}^M \hat{w}_k(t)} \sum_{k=1}^M 1_{\{\hat{\mathbf{V}}_k(t)=\mathbf{x}\}} \hat{w}_k(t).$$

Challenges

- **Sample Degeneracy Issue:** When the number of hidden species is large ($\dim(\mathbf{X}) \gg 1$), most importance weights $\hat{w}_k$ collapse to near-zero values, resulting in unreliable posterior estimates.
- **Inaccurate estimation of tail probabilities.**

Our approach in (Ben Hammouda et al. 2025): Motivation



- ① employs a reduced variance particle filter for estimating the jump intensities of the projected process (obtained with Markovian Projection), and
- ② solves the filtering equations in a low-dimensional space.

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Revisiting the Filtering Problem

Setting: we are only interested in a **subset of the hidden process**

$$\mathbf{Z} = \begin{pmatrix} \mathbf{X}' \\ \mathbf{X}'' \\ \mathbf{Y} \end{pmatrix}$$

hidden species of interest
remaining hidden species
observed species

Task: estimate the **marginal** conditional distribution of $\mathbf{X}'$:

$$\pi'^{\mathbf{y}}(\mathbf{x}', t) := \mathbb{P}_{\mu}(\mathbf{X}'(t) = \mathbf{x}' \mid \mathbf{Y}(s) = \mathbf{y}(s), 0 \leq s \leq t).$$

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1. Marginalization of the distribution:

Solve the filtering problem for d -dimensional process $\mathbf{Z}$ and compute

$$\pi'^{\mathbf{y}}(\mathbf{x}', t) = \sum_{\mathbf{x}''} \pi^{\mathbf{y}}(\mathbf{x}', \mathbf{x}'', t).$$

$\implies$ **curse of dimensionality**

Revisiting the Filtering Problem

Setting: we are only interested in a **subset of the hidden process**

$$\mathbf{Z} = \begin{pmatrix} \mathbf{X}' \\ \mathbf{X}'' \\ \mathbf{Y} \end{pmatrix}$$

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 remaining hidden species
 observed species

Task: estimate the **marginal** conditional distribution of $\mathbf{X}'$:

$$\pi'^{\mathbf{y}}(\mathbf{x}', t) := \mathbb{P}_{\mu}(\mathbf{X}'(t) = \mathbf{x}' \mid \mathbf{Y}(s) = \mathbf{y}(s), 0 \leq s \leq t).$$

2. Marginalization of the process:

Try to solve the filtering problem for d' -dimensional process

$$\mathbf{Z}' := \begin{pmatrix} \mathbf{X}' \\ \mathbf{Y} \end{pmatrix}.$$

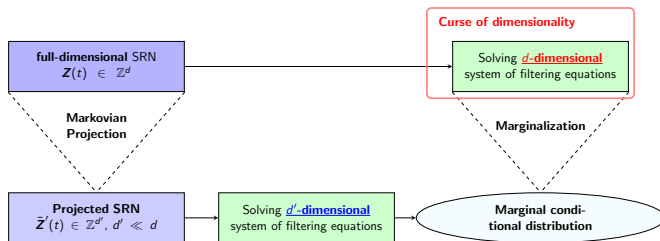
Challenge: $\mathbf{Z}'$ is not a Markov process.

Markovian Projection for Filtering

Challenge: Z' is a non-Markovian process.

Solution: construct a Markov process $\bar{Z}'$ that mimics the distribution of Z' .

$$Z = \begin{bmatrix} X' \\ X'' \\ Y \end{bmatrix} \begin{matrix} \rightarrow \begin{bmatrix} X' \\ Y \end{bmatrix} = Z' \\ \rightarrow \begin{bmatrix} X'' \end{bmatrix} = Z'' \end{matrix}$$



Standard Markovian Projection for Filtering

- Recall: $\mathbf{Z}(t) = \mathbf{Z}_0 + \sum_{j=1}^J Y_j \left(\int_0^t a_j(\mathbf{Z}(\tau)) d\tau \right) \cdot \boldsymbol{\nu}_j$
- Challenge:** $\mathbf{Z}'$ is not a Markov process.
- Solution:** construct a Markov process $\bar{\mathbf{Z}}'$ that mimics the distribution of $\mathbf{Z}'$.

Theorem (Ben Hammouda et al. 2024)

Let $\mathbf{Z}(t) = (\mathbf{Z}'(t), \mathbf{Z}''(t))^T$ be a non-explosive SRN (Assumption A.1). Define a d' -dimensional stochastic process $\bar{\mathbf{Z}}'$ via independent Poisson processes $\bar{R}_1, \dots, \bar{R}_J$ as

$$\bar{\mathbf{Z}}'(t) = \bar{\mathbf{Z}}'(0) + \sum_{j=1}^J \bar{R}_j \left(\int_0^t \bar{a}_j(\bar{\mathbf{Z}}'(s), s) ds \right) \boldsymbol{\nu}'_j \quad t \in [0, T],$$

where $\boldsymbol{\nu}'_j$ are stoichiometric vectors of $\mathbf{Z}'$, $\bar{\mathbf{Z}}'(0) \stackrel{d}{=} \mathbf{Z}'(0)$, and

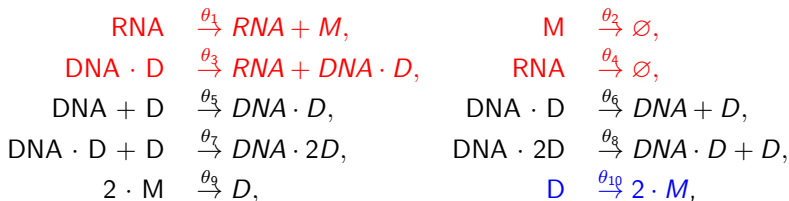
$$\bar{a}_j(\bar{\mathbf{Z}}', t) := \mathbb{E} [a_j(\mathbf{Z}(t)) \mid \mathbf{Z}'(t) = \bar{\mathbf{Z}}'] , \quad \bar{\mathbf{Z}}' \in \mathbb{Z}_{\geq 0}^{d'}$$

Then $\bar{\mathbf{Z}}'(t)$ has the same distribution as $\mathbf{Z}'(t)$ for all $t \in [0, T]$.

Markovian Projection: Example

Goutsias's model of regulated transcription ($d=6$, $J=10$) (Goutsias 2005)

$\mathbf{Z} := (M, D, RNA, DNA, DNA \cdot D(0), DNA \cdot 2D)$ with projection $\mathbf{Z}' = \mathbf{Z}_2$
(projection to molecule type D)

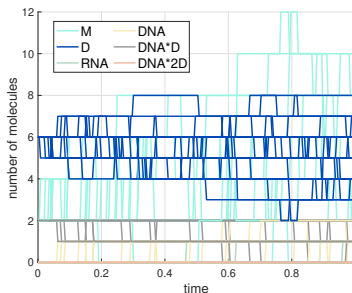


$$\bar{a}_{10}(t, \bar{\mathbf{z}}') = \mathbb{E} [\theta_{10} \mathbf{Z}_2(t) \mid \mathbf{Z}_2(t) = \bar{\mathbf{z}}'; \mathbf{Z}(0) = \mathbf{z}_0] = \theta_{10} \bar{\mathbf{z}}'$$

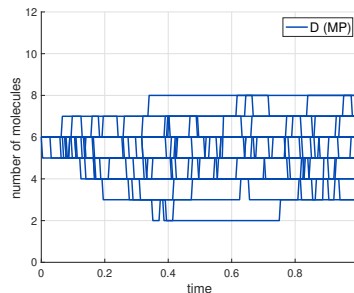
Reaction $j = 10$ is given in closed form and reactions $j = 1, 2, 3, 4$ do not occur in the projected process since $\nu'_j = 0$.

Markovian Projection: Numerical Experiments

Full-dimensional SRN (TL)



MP-SRN (TL)



30 sample paths generated with TL³ with step size $\Delta t = 2^{-8}$.

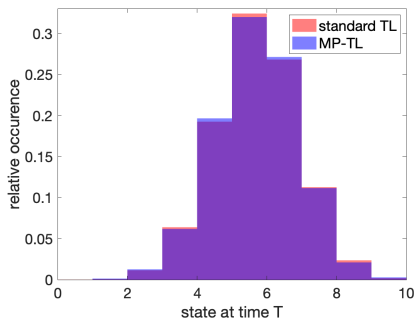
³Tau-Leap (TL) is the forward Euler approximation of Kurtz's random time-change representation

Markovian Projection: Numerical Illustration

Goutsias's model of regulated transcription

$(d=6, J=10, \bar{d}=1)$

$$P(\mathbf{x}) = x_2$$



Relative occurrences of states at final time T with 10^4 sample paths comparing the TL estimate of $P(\mathbf{X}(t)) |_{\{\mathbf{x}_0=\mathbf{x}_0\}}$ and the MP estimate of $\bar{\mathbf{S}}(T) |_{\{\mathbf{x}_0=\mathbf{x}_0\}}$.

Standard Markovian Projection for Filtering

Step 1: Construct projected SRN $\bar{\mathbf{Z}}'$ and estimate its propensities $\bar{a}_j$ using MC or L^2 regression.

$$\mathbf{Z} = \begin{pmatrix} \mathbf{x}' \\ \mathbf{x}'' \\ \mathbf{y} \end{pmatrix} \xrightarrow{\text{MP}} \bar{\mathbf{Z}}' = \begin{pmatrix} \bar{\mathbf{x}}' \\ \bar{\mathbf{y}} \end{pmatrix}$$

Step 2: Apply FFSP to solve the filtering problem for $\bar{\mathbf{Z}}'$ with dimension $d' \ll d$:

$$\bar{\pi}'^{\mathbf{y}}(\mathbf{x}', t) := \mathbb{P} \left(\bar{\mathbf{X}}'(t) = \mathbf{x}' \mid \bar{\mathbf{Y}}(s) = \mathbf{y}(s), 0 \leq s \leq t \right).$$

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Note

The MP theorem only states $\mathbf{Z}'(t) \stackrel{d}{=} \bar{\mathbf{Z}}'(t)$ for each t , which implies

$$\mathbb{P}(\mathbf{X}'(t) = \mathbf{x}' \mid \mathbf{Y}(t) = \mathbf{y}(t)) = \mathbb{P}(\bar{\mathbf{X}}'(t) = \mathbf{x}' \mid \bar{\mathbf{Y}}(t) = \mathbf{y}(t)).$$

But there is no guarantee that $\bar{\pi}'^{\mathbf{y}}$ will be equal to $\pi'^{\mathbf{y}}$.

⚠ **Recall:** $\pi'^{\mathbf{y}}(\mathbf{x}', t) := \mathbb{P} \left(\mathbf{X}'(t) = \mathbf{x}' \mid \mathbf{Y}(s) = \mathbf{y}(s), 0 \leq s \leq t \right).$

Filtered Markovian Projection (FMP)

Recall: for the MP SRN $\bar{\mathbf{Z}}'$ we have $\bar{\pi}'^{\mathbf{y}} \neq \pi'^{\mathbf{y}}$

Idea: construct another SRN that preserves the **conditional** distribution

Theorem (Ben Hammouda et al. 2025)

Let $\mathbf{Z} = (\mathbf{X}', \mathbf{X}'', \mathbf{Y})^\top$ be a non-explosive d -dimensional SRN (Assumption A.1) and let $\mathbf{Z}'(t) := (\mathbf{X}'(t), \mathbf{Y}(t)) \in \mathbb{Z}_{\geq 0}^{d'}$. Define a d' -dimensional stochastic process $\tilde{\mathbf{Z}}'(t) = (\tilde{\mathbf{X}}'(t), \tilde{\mathbf{Y}}(t))$ via independent Poisson processes $\tilde{R}_1, \dots, \tilde{R}_J$ as

$$\tilde{\mathbf{Z}}'(t) = \tilde{\mathbf{Z}}'(0) + \sum_{j=1}^J \tilde{R}_j \left(\int_0^t \tilde{a}_j(\tilde{\mathbf{Z}}'(s), s) ds \right) \boldsymbol{\nu}'_j, \quad t \in [0, T]$$

where $\boldsymbol{\nu}'_j$ are stoichiometric vectors of $\mathbf{Z}'$, $\tilde{\mathbf{Z}}'(0) \stackrel{d}{=} \mathbf{Z}'(0)$ and

$$\tilde{a}_j(\tilde{\mathbf{Z}}', t) := \mathbb{E}[a_j(\mathbf{Z}(t)) | \mathbf{Z}'(t) = \tilde{\mathbf{Z}}', \mathbf{Y}(s) = \mathbf{y}(s), s \leq t].$$

Then $\tilde{\mathbf{Z}}'(t)$ has the same **conditional** on $\{\tilde{\mathbf{Y}}(s) = \mathbf{y}(s), s \leq t\}$ distribution as $\mathbf{Z}'(t)$ **conditional** on $\{\mathbf{Y}(s) = \mathbf{y}(s), s \leq t\}$ for any $t \in [0, T]$.

FMP for Filtering

Step 1: Construct projected SRN $\tilde{\mathbf{z}}'$

$$\mathbf{Z} = \begin{pmatrix} \mathbf{x}' \\ \mathbf{x}'' \\ \mathbf{Y} \end{pmatrix} \xrightarrow{\text{FMP}} \tilde{\mathbf{z}}' = \begin{pmatrix} \tilde{\mathbf{x}}' \\ \tilde{\mathbf{Y}} \end{pmatrix}$$

and estimate its propensities with [Particle Filter](#)

$$\tilde{a}_j(\tilde{\mathbf{z}}', t) = \mathbb{E} [a_j(\mathbf{Z}(t)) \mid \mathbf{Z}'(t) = \tilde{\mathbf{z}}', \mathbf{Y}(s) = \mathbf{y}(s), 0 \leq s \leq t]$$

Step 2: Apply FFSP to solve the filtering problem for $\tilde{\mathbf{z}}'$ with dimension $d' \ll d$:

$$\tilde{\pi}'^{\mathbf{y}}(\mathbf{x}', t) := \mathbb{P} [\tilde{\mathbf{X}}'(t) = \mathbf{x}' \mid \tilde{\mathbf{Y}}(s) = \mathbf{y}(s), 0 \leq s \leq t].$$

FMP for Filtering

Step 1: Construct projected SRN $\tilde{\mathbf{z}}'$

$$\mathbf{Z} = \begin{pmatrix} \mathbf{x}' \\ \mathbf{x}'' \\ \mathbf{y} \end{pmatrix} \xrightarrow{\text{FMP}} \tilde{\mathbf{z}}' = \begin{pmatrix} \tilde{\mathbf{x}}' \\ \tilde{\mathbf{y}} \end{pmatrix}$$

and estimate its propensities with **Particle Filter**

$$\tilde{a}_j(\tilde{\mathbf{z}}', t) = \mathbb{E} [a_j(\mathbf{Z}(t)) \mid \mathbf{z}'(t) = \tilde{\mathbf{z}}', \mathbf{Y}(s) = \mathbf{y}(s), 0 \leq s \leq t]$$

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Note

The FMP theorem guarantees $\pi'^{\mathbf{y}} = \tilde{\pi}'^{\mathbf{y}}$:

$$\mathbb{P} [\mathbf{x}'(t) = \mathbf{x}' \mid \mathbf{Y}(s) = \mathbf{y}(s), s \leq t] = \mathbb{P} [\tilde{\mathbf{x}}'(t) = \mathbf{x}' \mid \tilde{\mathbf{y}}(s) = \mathbf{y}(s), s \leq t].$$

Error analysis

- We have an approximation $\tilde{\pi}_{FFSP}^{M,\Delta t} \approx \pi$ with the following error sources
 - **Projection error** due to the approximation of $\tilde{a}_j$ with M particles.
 - **Truncation error** due to the state space truncation in FFSP method.
 - **ODE solver error** due to the discretization of FFSP system with time step Δt .
 - The **model reduction error** is zero (Theorem 2).
- The total error is decomposed as follows:

$$\begin{aligned}
 \left| \pi(\mathbf{x}', t) - \tilde{\pi}_{FFSP}^{M,\Delta t}(\mathbf{x}', t) \right| &\leq \underbrace{\left| \pi(\mathbf{x}', t) - \tilde{\pi}(\mathbf{x}', t) \right|}_{\text{Model reduction error}} + \underbrace{\left| \tilde{\pi}(\mathbf{x}', t) - \tilde{\pi}^M(\mathbf{x}', t) \right|}_{\text{Projection error}} \\
 &\quad + \underbrace{\left| \tilde{\pi}^M(\mathbf{x}', t) - \tilde{\pi}_{FFSP}^M(\mathbf{x}', t) \right|}_{\text{Truncation error}} + \underbrace{\left| \tilde{\pi}_{FFSP}^M(\mathbf{x}', t) - \tilde{\pi}_{FFSP}^{M,\Delta t}(\mathbf{x}', t) \right|}_{\text{ODE solver error}}
 \end{aligned}$$

Error analysis

In practice, we have an approximation $\tilde{\pi}_{FFSP}^{M,\Delta t} \approx \pi$ with the following error sources

- **Projection error** due to the approximation of $\tilde{a}_j$ with M particles.
- **Truncation error** due to the state space truncation in FFSP method.
- **ODE solver error** due to the discretization of FFSP system with time step Δt .

Theorem (Ben Hammouda et al. 2025)

Under some assumptions (see (Ben Hammouda et al. 2025) for more details), we show

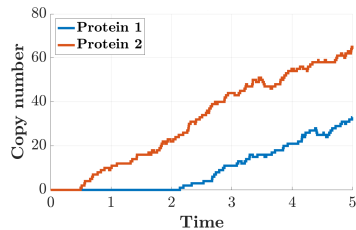
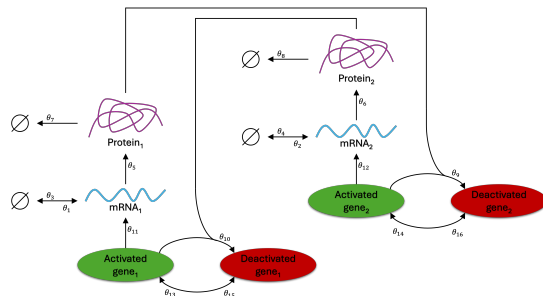
$$\mathbb{E} \left[\left\| \pi(\mathbf{x}', t) - \tilde{\pi}_{FFSP}^{M,\Delta t}(\mathbf{x}', t) \right\|_1 \right] \leq \frac{C_1}{\sqrt{M}} + \mathcal{E}_{FFSP} + C_2 \Delta t^p$$

- ✓ We can control $\mathcal{E}_{FFSP}$ and $C_2 \Delta t^p$ with little cost since the state space of projected SRN is relatively small (potentially one-dimensional).

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Example 1: Bistable Gene Switch Network

Consider the following SRN (Gupta et al. 2021) with: 2 observed and 6 hidden species, and 16 reactions.

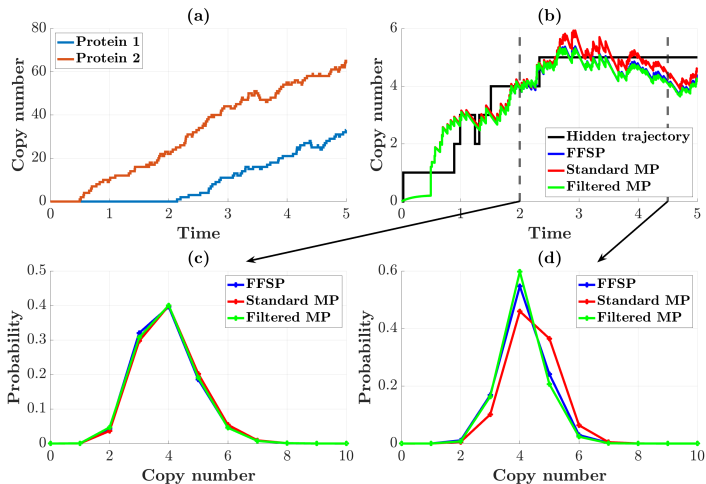


Observed trajectory of Protein₁ and Protein₂.

Task: Estimate the distribution of the *mRNA₂* copy number, i.e.,

$$\pi'(x', t) := \mathbb{P} \left\{ mRNA_2(t) \mid \begin{bmatrix} protein_1(s) = y_1(s) \\ protein_2(s) = y_2(s) \end{bmatrix}, 0 \leq s \leq t \right\}$$

Bistable Gene Switch Network: Numerical Results



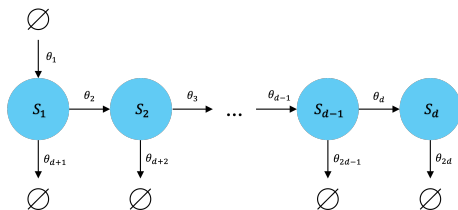
(a) Observed trajectories, (b) Estimated conditional mean, (c)-(d) Estimated conditional distribution at times $t = 2$ and $t = 4.5$.

Bistable Gene Switch Network: Numerical Results

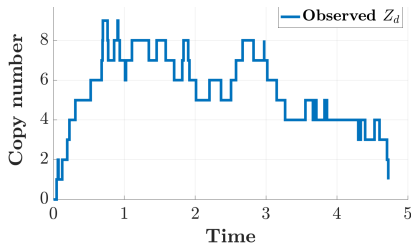
	Dimension of hidden space	# of hidden states	CPU time in seconds
FFSP (full model)	6	15 376	846.4 s
MP-FFSP (projected model)	1	31	5.1 s
FMP-FFSP (projected model)	1	31	5.7 s

The Computational complexity of solving the filtering equations with FFSP for full and projected models.

Example 2: Linear Cascade Model (Gupta et al. 2021)



Reactions in linear cascade SRN



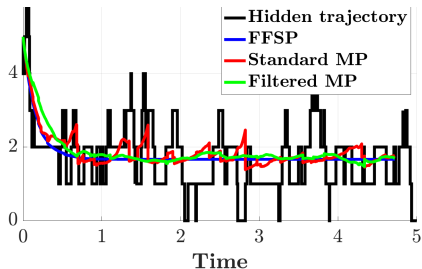
Observed trajectory of S_d .

Task: Estimate the the distribution of S_1 given that S_d is observed, i.e.,

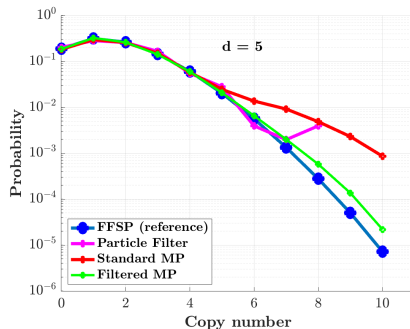
$$\pi'(x', t) = \mathbb{P}\{Z_1(t) \mid Z_d(s) = y(s), 0 \leq s \leq t\}$$

Linear Cascade Model: Numerical Results

$$d = 5$$

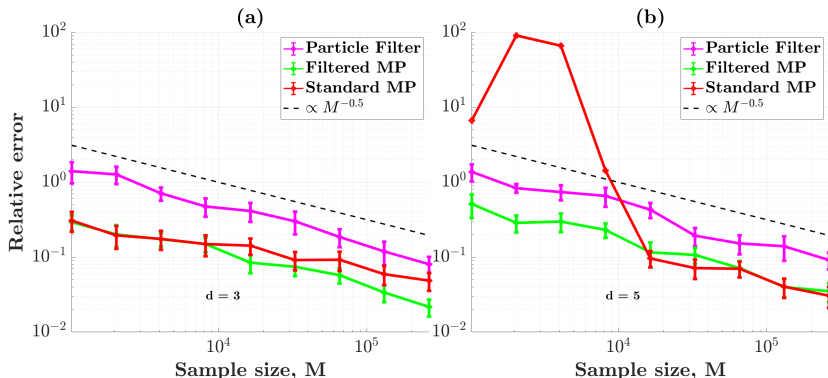


Estimated conditional mean
 $\mathbb{E}[Z_1(t) | Z_d(s) = y(s), s \leq t]$



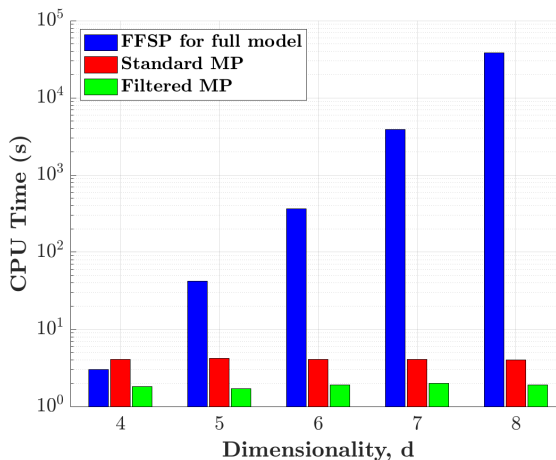
Estimated conditional distribution
 $\mathbb{P}\{Z_1(T) | Z_d(s) = y(s), s \leq T\}$
in log scale

Linear Cascade Model: Numerical Results



Relative error for estimating $\mathbb{P}[Z_1(T) \geq 8 \mid Z_d(s) = y(s), 0 \leq s \leq T] \approx 10^{-4}$. (a) the 3-dimensional model. (b) the 5-dimensional model.

Linear Cascade Model: Numerical Results

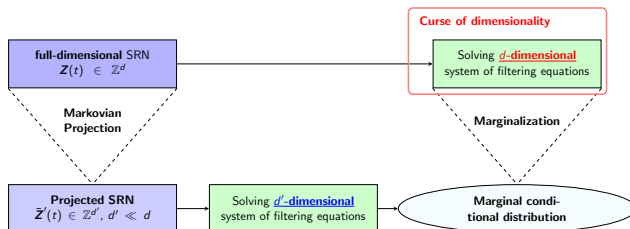


Computational cost versus dimension for solving the filtering problem

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Conclusion

- Existing filtering methods for SRNs are not efficient in the **high dimensional** setting.
- Filtered Markovian Projection** (FMP) guarantees convergence, improves accuracy and efficiency compared to particle filters, and mitigates the curse of dimensionality faced by FFSP when applied to the full model:
 - Employs a reduced-variance particle filter to estimate the jump intensities of the projected process (obtained via Markovian projection),
 - Solves the filtering equations in a low-dimensional space.



Work in Progress and Future Directions

- Adapt FMP from (Ben Hammouda et al. 2025) to settings with **noisy and discrete observations**.
- Extend FMP to include **parameter estimation** (via particle filters or other Bayesian inference methods) $\Rightarrow$ higher gains expected, but more challenging.
- Combine MP with particle filters to reduce sample degeneracy and improve tail accuracy (Djurić et al. 2013; Snyder et al. 2008), in particular, using filtered MP as Control Variate for particle filters.
- Explore **Koopman-based decompositions** to separate time scales (Li et al. 2023) $\Rightarrow$ reduce dimensionality w.r.t. number of reactions.

Thank you for your attention!

- ① C. Ben Hammouda, M. Chupin, S. Münker, and R. Tempone. “Filtered Markovian Projection: Dimensionality Reduction in Filtering for Stochastic Reaction Networks”. In: *arXiv preprint arXiv:2502.07918* (2025)
Accompanying code can be found here:
[Git repository: Markovian-Projection-in-filtering-for-SRNs](#)
- ② C. Ben Hammouda, N. Ben Rached, R. Tempone, and S. Wiechert. “Automated importance sampling via optimal control for stochastic reaction networks: A Markovian projection–based approach”. In: *Journal of Computational and Applied Mathematics* 446 (2024), p. 115853
- ③ C. Ben Hammouda, N. Ben Rached, R. Tempone, and S. Wiechert. “Learning-based importance sampling via stochastic optimal control for stochastic reaction networks”. In: *Statistics and Computing* 33.3 (2023), p. 58

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